



Stockholms  
universitet

**OBS!** Läs noga igenom anvisningarna i tentamen, t.ex. hur du ska skriva svaren.  
Det är ditt ansvar som student att följa de anvisningar som ges.

**NOTE!** Read the examination instructions carefully, e.g. how to write the answers.  
It is your responsibility as a student to follow the given instructions.

Skriv din anonymiseringskod och dagens datum på allt material du lämnar in.  
(Enter your anonymization code and today's date on all submitted materials)

Anonymiseringskod (Anonymization code)	3	1	1	-	0	0	0	7	-	R	L	M
Datum (Date YYYY-MM-DD)	2024-01-04							Plats nr. (Seat No.)	50			

Kurs/Kurskod (Course/Course code)	ST721A
Kursmoment (Course component)	Sannolikhetssteori

Fylls i av tentamensvärd (To be filled in by invigilator)

Direkt i skrivning: (kryss)		Svarsblankett: (kryss)		Lösa svarsblad: (antal)	10
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Lämnat in blankt: (kryss)		Dator: (kryss)	
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Inlämningstid: 13 : 25 Signatur tentamensvärd: Li Selen

Fylls i av lärare/examinator (To be filled in by teacher/examinator)

Betyg:	D	Poäng:	63
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Signatur rättande lärare/examinator: AA

1	2	3	4	5	6	7	8	9	21
10	8	4	12	11	12	6	-	-	(63)



$$f_{xy}(x,y) = \begin{cases} 6xy^2, & 0 < x < 1, 0 < y < 1 \\ 0, & \text{otherwise} \end{cases}$$

Uppg.nr.: (Task no.)

1

Lärarens kommentar: (Teacher's note)

*$f(x,y) \geq 0$ ?*

$$(a) \int_0^1 \int_0^1 6xy^2 dx dy = \int_0^1 [6y^2 \frac{x^2}{2}]_0^1 dy = \int_0^1 3y^2 dy = [y^3]_0^1 = 1 \text{ ok}$$

$$(b) P(X+Y \geq 0.9) = P(X \geq 0.9 - Y) =$$

$$\int_0^1 \int_{0.9-y}^1 6xy^2 dx dy = \int_0^1 6y^2 [\frac{x^2}{2}]_{0.9-y}^1 dy = \int_0^1 6y^2 [1 - \frac{(0.9-y)^2}{2}] dy$$

$$= 0.94 \text{ ok}$$

$$(c) \int_0^1 \int_{1/2}^1 6xy^2 dx dy - \int_0^1 \int_0^{1/2} 6xy^2 dx dy = \int_0^1 6y^2 [\frac{x^2}{2}]_{1/2}^1 dy - \int_0^1 6y^2 [\frac{x^2}{2}]_0^{1/2} dy$$

$$= \int_0^1 6y^2 (\frac{1}{2} - \frac{1}{8}) dy - \int_0^1 6y^2 (\frac{1}{8}) dy = \int_0^1 \frac{9}{4} y^2 dy - \int_0^1 \frac{3}{4} y^2 dy =$$

$$= [\frac{9}{4} \frac{y^3}{3}]_0^1 - [\frac{3}{4} \frac{y^3}{3}]_0^1 = \frac{9}{12} - \frac{3}{12} = \frac{6}{12} = \frac{1}{2}$$

We get that the probability of  $x$  is  $\frac{1}{2} < x < 1$  is 50% - units greater than if  $x$  is  $0 < x < \frac{1}{2}$ . So  $P(\frac{1}{5} < x < 1) = 75\%$  and  $P(0 < x < \frac{1}{2}) = 25\%$ .

*ok for the calculation but partial credit on discussion. ok\_*

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2)  $f_{xy}(x,y) = \begin{cases} x+y, & 0 < x < 1, 0 < y < 1 \\ 0, & \text{otherwise} \end{cases}$

Uppg.nr.: (Task no.)

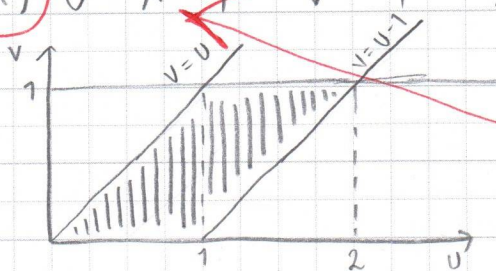
2

Lärarens kommentar: (Teacher's note)

$$\iint_0^1 (x+y) dx dy = \int_0^1 [x^2/2 + xy]_0^1 dy = \int_0^1 [1/2 + y] dy = [y/2 + y^2/2]_0^1$$

$$= 1/2 + 1/2 = 1 \quad \text{ok}$$

a)  $U = X + Y \quad V = Y \quad X = U - V \quad Y = V$



$$0 < U - V < 1, \quad 0 < V < 1 \\ V < U < V + 1$$

$$f_{uv}(u,v) = f_{xy}(u-v, v) \cdot |J|$$

$$\begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} 1 & -1 \\ 0 & 1 \end{vmatrix}$$

$$J = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} 1 & -1 \\ 0 & 1 \end{vmatrix} = 1 \cdot 1 - 0 \cdot (-1) = 1$$

$$f_{uv}(u,v) = (u-v) + v = u, \quad 0 < v < 1, \quad v < u < v + 1$$

$$\int_0^u u dv = [uv]_0^u = u^2, \quad 0 < u < 1$$

$$f(u) = \int_{u-1}^1 u dv = [uv]_{u-1}^1 = u - u(u-1) = 2u - u^2, \quad 1 \leq u < 2$$

$$\begin{cases} (x+y)^2, & 0 < x+y < 1 \\ 2(x+y) - (x+y)^2, & 1 \leq x+y < 2 \end{cases}$$

ok but better to write in terms of  $U$

$$f_{xy}(x+y) = \begin{cases} (x+y)^2, & 0 < x+y < 1 \\ 2(x+y) - (x+y)^2, & 1 \leq x+y < 2 \\ 0, & \text{otherwise} \end{cases}$$

Tentamensservice/SU/typ\_b

b)  $E(X+Y) = \int_0^1 (x+y)^3 d(x+y) + \int_1^2 (2(x+y) - (x+y)^3) d(x+y) = \int_0^1 u^3 du + \int_1^2 u^2 \cdot (2u - u^2) du$

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cont 2. b)

$$= \left[ \frac{(x+y)^4}{4} \right]_0^1 + \left[ \frac{2(x+y)^3}{3} - \frac{(x+y)^4}{4} \right]_1^2 = \frac{1}{4} + \frac{16}{3} - \frac{16}{4} - \frac{2}{3} + 2$$

+  $\frac{1}{4} = \frac{7}{6}$  or (mean) (median)

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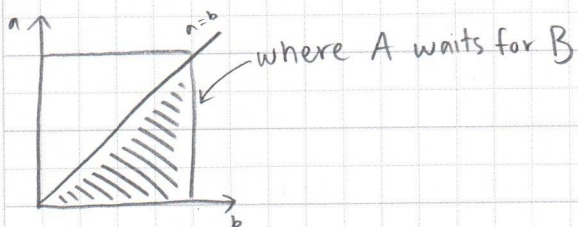




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3  $A, B \stackrel{i.i.d}{\sim} U[13, 18]$  *OK*  
 $f_A(a) = f_B(b) = 1/5, 13 \leq a, b \leq 18$  *OK*

Because of independency:  
 $f_{AB}(a, b) = f_A(a) \cdot f_B(b) = 1/25, 13 \leq a, b \leq 18$  *OK*



distribution?

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Lärarens kommentar:  
(Teacher's note)

$C =$  "waiting time"

$$\text{distr. of } C = \int_{13}^{18} \int_b^{18} \frac{1}{25} da db = \int_{13}^{18} \left[ \frac{a}{25} \right]_b^{18} db = \int_{13}^{18} \frac{18}{25} - \frac{b}{25} db =$$

$$= \left[ \frac{18b}{25} - \frac{b^2}{50} \right]_{13}^{18} = \left( \frac{18^2}{25} - \frac{18^2}{50} \right) - \left( \frac{18 \cdot 13}{25} - \frac{13^2}{50} \right) =$$

$$= \frac{1}{2}$$

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4)  $X, Y \stackrel{iid}{\sim} N(1, 4)$

$U = X + Y \quad V = X - Y$

a)  $U + V = (X + Y) + (X - Y) = 2X$

$X = \frac{U + V}{2}$

$U - V = (X + Y) - (X - Y) = 2Y$

$Y = \frac{U - V}{2}$

When independent  $f_{xy}(x, y) = g_x(x) \cdot h_y(y)$

$g_x(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-1)^2}{8}} \quad -\infty < x < \infty$

$h_y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-1)^2}{8}} \quad -\infty < y < \infty$

$f_{xy}(x, y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-1)^2}{8}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-1)^2}{8}} \quad -\infty < x, y < \infty$

$f_{uv}(u, v) = f_{xy}\left(\frac{u+v}{2}, \frac{u-v}{2}\right) \cdot |J|$

$|J| = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{vmatrix}$

$= \frac{1}{2} \cdot (-\frac{1}{2}) - \frac{1}{2} \cdot \frac{1}{2} = -\frac{1}{4} - \frac{1}{4} = -\frac{1}{2}$

$f_{uv}(u, v) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{8} \left( \left(\frac{u+v}{2} - 1\right)^2 + \left(\frac{u-v}{2} - 1\right)^2 \right)} \cdot \frac{1}{2}$

b) We can rewrite  $f_{uv}(u, v) = g_u(u) \cdot h_v(v)$

$f_{uv}(u, v) = \underbrace{\frac{1}{\sqrt{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{8} \left(\frac{u^2}{2} - 2u\right)}}_{g_u(u)} \cdot \underbrace{\frac{1}{\sqrt{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{8} \left(\frac{v^2}{2} + 2\right)}}_{h_v(v)}$

Because  $U$  and  $V$  are independent,  $cov(U, V) = 0$

and  $corr(U, V) = \frac{cov(U, V)}{\sigma_u \sigma_v} = \frac{0}{2 \cdot 2} = 0$

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5)  $f_{xy}(x,y) = 1, 0 < y < 1, y < x < y+1$

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Lärarens kommentar:  
(Teacher's note)

$$\int_0^1 \int_y^{y+1} 1 dx dy = \int_0^1 [x]_y^{y+1} dy = \int_0^1 1 dy = [y]_0^1 = 1 \quad \text{OK}$$

$Y \sim U[0,1] \quad f_Y(y) = 1, 0 < y < 1$

Draw domain?

a)  $\int_0^x 1 dy = [y]_0^x = x, 0 < x < 1$

OK

$f_X(x) = \int_{x-1}^1 1 dy = [y]_{x-1}^1 = 2-x, 1 \leq x < 2$

$$f_X(x) = \begin{cases} x, & 0 < x < 1 \\ 2-x, & 1 \leq x < 2 \end{cases}$$

OK

$f_X(x) = 0, \text{ otherwise}$

$5Y \sim U[0,5]$

$f_{5Y}(5y) = 1/5, 0 < y < 5$

OK

$$E(X) = \int_0^1 x^2 dx + \int_1^2 (2x - x^2) dx = \left[ \frac{x^3}{3} \right]_0^1 + \left[ x^2 - \frac{x^3}{3} \right]_1^2 = \frac{1}{3} +$$

$$+ 4 - \frac{8}{3} - 1 + \frac{1}{3} = 1$$

OK

$$E(X^2) = \int_0^1 x^3 dx + \int_1^2 (2x^2 - x^3) dx = \left[ \frac{x^4}{4} \right]_0^1 + \left[ \frac{2x^3}{3} - \frac{x^4}{4} \right]_1^2 = \frac{1}{4} +$$

$$+ \frac{16}{3} - 4 - \frac{2}{3} + \frac{1}{4} = \frac{7}{6}$$

$\text{Var}(X) = E(X^2) - (E(X))^2 = \frac{7}{6} - 1 = \frac{1}{6}$

OK

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Lärarens kommentar:  
(Teacher's note)

cont. 5. a)

$$E(5Y) = 5 \cdot E(Y) = 5 \cdot 1/2 = 5/2 \text{ ok}$$

$$\text{Var}(5Y) = 25 \cdot \text{Var}(Y) = 25 \cdot 1/12 = 25/12 \text{ ok } (+)$$

b)  $\text{Corr}(X, 5Y) = \text{Corr}(X, Y)$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$E(XY) = \int_0^1 \int_0^{y+1} xy \, dx \, dy = \int_0^1 y \left[ \frac{x^2}{2} \right]_0^{y+1} dy = \int_0^1 y^2 + \frac{y}{2} dy =$$

$$= \left[ \frac{y^3}{3} + \frac{y^2}{4} \right]_0^1 = \frac{1}{3} + \frac{1}{4} = \frac{7}{12}$$

$$\text{Cov}(X, Y) = \frac{7}{12} - 1 \cdot \frac{1}{2} = \frac{1}{12}$$

$$\text{Corr}(X, Y) = \frac{\frac{1}{12}}{\sqrt{\frac{1}{6}} \cdot \sqrt{\frac{1}{12}}} = \frac{\sqrt{\frac{1}{144}}}{\sqrt{\frac{1}{72}}} = \sqrt{\frac{1}{144}} \cdot \sqrt{\frac{72}{1}} = \sqrt{\frac{72}{144}}$$

$$\frac{1}{\cancel{6}} \cdot \frac{1}{2} \cdot \frac{\sqrt{6} \cdot \sqrt{6} \cdot \sqrt{2}}{\sqrt{12}} = \frac{1}{\sqrt{2}} \text{ (+)}$$

otherwise, very good.

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6  $X = \text{minimum}$        $Y = \text{maximum}$

$y \backslash x$	1	2	3	4	5	6	$f_Y(y)$
1	$1/36$	0	0	0	0	0	$1/36$
2	$2/36$	$1/36$	0	0	0	0	$3/36$
3	$2/36$	$2/36$	$1/36$	0	0	0	$5/36$
4	$2/36$	$2/36$	$2/36$	$1/36$	0	0	$7/36$
5	$2/36$	$2/36$	$2/36$	$2/36$	$1/36$	0	$9/36$
6	$2/36$	$2/36$	$2/36$	$2/36$	$2/36$	$1/36$	$11/36$
$f_X(x)$	$11/36$	$9/36$	$7/36$	$5/36$	$3/36$	$1/36$	1

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Lärarens kommentar:  
(Teacher's note)

$$E(X) = (1 \cdot 11 + 2 \cdot 9 + 3 \cdot 7 + 4 \cdot 5 + 5 \cdot 3 + 6 \cdot 1) / 36 = 91/36$$

$$E(X^2) = (1^2 \cdot 11 + 2^2 \cdot 9 + 3^2 \cdot 7 + 4^2 \cdot 5 + 5^2 \cdot 3 + 6^2 \cdot 1) / 36 = 307/36$$

$$\text{Var}(X) = 307/36 - (91/36)^2 = 2555/1296$$

$$E(Y) = (1 \cdot 1 + 2 \cdot 3 + 3 \cdot 5 + 4 \cdot 7 + 5 \cdot 9 + 6 \cdot 11) / 36 = 167/36$$

$$E(Y^2) = (1^2 \cdot 1 + 2^2 \cdot 3 + 3^2 \cdot 5 + 4^2 \cdot 7 + 5^2 \cdot 9 + 6^2 \cdot 11) / 36 = 791/36$$

$$\text{Var}(Y) = 791/36 - (167/36)^2 = 2555/1296$$

$$E(XY) = (1 \cdot 1 \cdot 1 + 1 \cdot 2 \cdot 2 + 1 \cdot 3 \cdot 2 + 1 \cdot 4 \cdot 2 + 1 \cdot 5 \cdot 2 + 1 \cdot 6 \cdot 2 + 2 \cdot 2 \cdot 1 + 2 \cdot 3 \cdot 2 + 2 \cdot 4 \cdot 2 + 2 \cdot 5 \cdot 2 + 2 \cdot 6 \cdot 2 + 3 \cdot 3 \cdot 1 + 3 \cdot 4 \cdot 2 + 3 \cdot 5 \cdot 2 + 3 \cdot 6 \cdot 2 + 4 \cdot 4 \cdot 1 + 4 \cdot 5 \cdot 2 + 4 \cdot 6 \cdot 2 + 5 \cdot 5 \cdot 1 + 5 \cdot 6 \cdot 2 + 6 \cdot 6 \cdot 1) / 36 = 49/4$$

$$\text{Cov}(X, Y) = 49/4 - 91/36 \cdot 167/36 = 1225/1296$$

$$\text{Corr}(X, Y) = \frac{\sqrt{2555/1296} \cdot \sqrt{2555/1296}}{2555/1296} = 35/73$$

We have a positive correlation around  $\sim 0.48$ , which corresponds to the value of the max and min with the two tosses.

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Lärarens kommentar:  
(Teacher's note)

7) a)  $f_x(x) = \frac{1}{2\beta} e^{-|x-\alpha|/\beta}$   $-\infty < x, \alpha < \infty, \beta > 0$

We know that double exponential is a pdf so we say  $\alpha = \mu$  and  $\beta = \sigma$  and get following:

$g_x(x) = \frac{1}{2\sigma} e^{-|x-\mu|/\sigma}$  and  $\int_{-\infty}^{\infty} \frac{1}{2\sigma} e^{-|x-\mu|/\sigma} dx = 1$   
 $-\infty < x < \infty, -\infty < \mu < \infty, \sigma > 0$

$M_x(t) = E(e^{tx})$

$E(e^{tx}) = \int_{-\infty}^{\infty} e^{tx} \frac{1}{2\beta} e^{-|x-\alpha|/\beta} dx = \int_{-\infty}^{\infty} \frac{1}{2\beta} e^{tx - |x-\alpha|/\beta} dx =$

$\frac{1 - (\beta t)^2}{e^{t\alpha}}$

b)  $X \sim \text{Gamma}(\alpha, \beta)$

$f_x(x) = \frac{1}{\Gamma(\alpha) \beta^\alpha} x^{\alpha-1} e^{-x/\beta}, 0 \leq x < \infty, \alpha, \beta > 0$

$M_x(t) = E(e^{tx}) = \int_0^{\infty} e^{tx} \frac{1}{\Gamma(\alpha) \beta^\alpha} x^{\alpha-1} e^{-x/\beta} dx$  ok

$tx - x/\beta = x(\beta t - 1/\beta) = -x(1 - \beta t/\beta) = -x/(\beta/(1 - \beta t))$

We know that:

$\int_0^{\infty} \frac{1}{\Gamma(\alpha) \beta^\alpha} x^{\alpha-1} e^{-x/\beta} dx = 1$

$\int_0^{\infty} x^{\alpha-1} e^{-x/\beta} dx = \Gamma(\alpha) \beta^\alpha$

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Lärarens kommentar:  
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cont. 7. b)

$$\frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^\infty x^{\alpha-1} e^{-x/(\beta/(1-\beta t))} dx =$$

$$\frac{1}{\Gamma(\alpha)\beta^\alpha} \Gamma(\alpha) \left(\frac{\beta}{1-\beta t}\right)^\alpha = \frac{1}{\beta^\alpha} \frac{\beta^\alpha}{(1-\beta t)^\alpha} = \left(\frac{1}{1-\beta t}\right)^\alpha$$

if  $t > 1/\beta$

$$E(X) = M'_x(0) = \frac{d}{dt} M_x(t) \Big|_{t=0} = \frac{d}{dt} \left(\frac{1}{1-\beta t}\right)^\alpha =$$

$$\frac{d}{dt} (1-\beta t)^{-\alpha} = (-\alpha) \cdot (1-\beta t)^{-\alpha-1} \cdot (-\beta) = \alpha \cdot \beta$$

$$E(X^2) = M''_x(0) = \frac{d^2}{dt^2} M_x(t) \Big|_{t=0} = \frac{d}{dt} \alpha \cdot \beta (1-\beta t)^{-\alpha-1} =$$
$$= \alpha \cdot (-\alpha-1) \cdot \beta \cdot (-\beta) (1-\beta t)^{-\alpha-2} = \alpha^2 \beta^2 + \alpha \beta^2$$

$$\text{Var}(X) = \alpha^2 \beta^2 + \alpha \beta^2 - (\alpha \cdot \beta)^2 = \alpha \beta^2$$

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## Regler i skrivsalen

- Följ tentamensvärds anvisningar.
- Väskor och ytterkläder ska placeras på anvisad plats.
- Placera ID-handling väl synlig på bordet framför dig.
- Ingen student får lämna skrivsalen under de första 30 minuterna.
- Endast en student i taget får besöka toaletten. Vid toalettbesök skriv ditt namn och klockslag på avsedd lista. Efter toalettbesöket ska du åter ange klockslag på listan.
- Elektronisk utrustning som mobiltelefon eller Smartwatch ska vara avstängd och placerad på anvisad plats.
- Under tentamen gäller tystnad – det är förbjudet att prata, eller på annat sätt kommunicera, med andra studenter under pågående tentamen.
- Innan tentamenshandlingarna lämnas in; skriv sidnummer, anonymiseringskod och datum på alla inlämnade papper.

Om något är oklart – fråga gärna tentamensvärden. Lycka till!

## Rules in the examination hall

- Follow the invigilator's instructions.
- Bags and outerwear must be placed at the designated place.
- Place your ID document clearly visible on the table in front of you.
- No student may leave the examination hall for the first 30 minutes.
- Only one student at a time may visit the toilet. Before visiting the toilet, write your name and time on the intended list. After the toilet visit, enter the time on the list again.
- Electronic equipment such as a mobile phone or Smartwatch must be switched off and placed at the designated place.
- During the exam, silence applies – you are not allowed to talk, or otherwise communicate, with other students during the exam.
- Before submitting the examination documents; remember to write the page number, anonymization code, and date on all papers.

Please do not hesitate to ask the invigilator if anything is unclear. Good luck!