

1 a) In order to test $H_0: \beta_1 = \beta_2 = \beta_3 = 0$ vs $H_1: \neg H_0$

I will first calculate R^2 .

$$R^2 = 1 - \frac{SSR}{SST} = 1 - \frac{5078,71318}{85260,44444} = 0,9404329497$$

$$n = 18$$

$$K = 3$$

$$F = \frac{R^2/K}{(1-R^2)/(n-K-1)} \sim F_{K, n-K-1} \text{ (if } H_0 \text{ is correct)}$$

$$F_{obs} = \frac{R^2/K}{(1-R^2)/(n-K-1)} = \frac{0,9404329497/3}{(1-0,9404329497)/(18-3-1)} = 73,67642139$$

We reject H_0 if $F_{obs} = 73,67642139 > F_{3,14} \Rightarrow c_{0,05} = 3,34$

Since $F_{obs} \approx 73 > F_{3,14} = 3,34$ we reject H_0 at 5% significance.

b) The test statistic will be:

$$F = \frac{(R_{ur}^2 - R_r^2)/q}{(1 - R_{ur}^2)/(n - k - 1)} \sim F_{q, n-k-1} \text{ (if } H_0 \text{ is correct)}$$

We have $R_r^2 = 0,9404329497$ and SST is the same since no new variables were added.

$$R_{ur}^2 = \frac{SSE_{ur}}{SST} = \frac{82707,71658195}{85260,44444} = 0,9700603501$$

$$n = 18$$

$$K = 5$$

$$q = 2$$

so $H_0: \beta_4 = \beta_5 = 0$ vs $H_1: \neg H_0$

$$F_{obs} = \frac{(0,9700603501 - 0,9404329497)/2}{(1 - 0,9700603501)/(18 - 5 - 2)} = 5,442638864$$

H_0 is rejected if $F_{obs} \approx 5,44 > F_{2,11} \Rightarrow c_{0,05} = 3,98$

Since $F_{obs} > F_{2,11}$ at 5% significance level H_0 is rejected.

0004-MXU

2. a) The dependent variable is total number of children ever born to a woman.

The interpretation of the coefficient on $y82$ is that the expected number of children ever born to a woman drops by 19,26076% if she gave birth to a child in the year 1982.

b) The fertility of black women is estimated to be 36,03475% higher than that of non black women holding other factors fixed.

c) Since the dispersion parameter $\sigma^2=1$ the standard deviation $\sigma=1$.

This indicates that the Poisson variance assumption is fulfilled and does not point to under- or overdispersion.

0004-MXU

$$3. \quad Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + u$$

In order to test $H_0: \beta_1 + \beta_2 = 1$ we first write it as $H_0: \beta_1 + \beta_2 - 1 = 0$. We then define $\theta_1 = \beta_1 + \beta_2 - 1$ and rewrite it as $\beta_1 = \theta_1 - \beta_2 + 1$.

Now we substitute this into the model which gives

$$Y = \beta_0 + (\theta_1 - \beta_2 + 1)X_1 + \beta_2 X_2 + u \Rightarrow$$

$$\Rightarrow Y = \beta_0 + \theta_1 X_1 - \beta_2 X_1 + X_1 + \beta_2 X_2 + u \Rightarrow$$

$$\Rightarrow Y = \beta_0 + \theta_1 X_1 + \beta_2 (X_2 - X_1) + X_1 + u$$

Since we have a X_1 without a parameter we subtract this on both sides which gives

$$Y - X_1 = \beta_0 + \theta_1 X_1 + \beta_2 (X_2 - X_1) + u$$

This model can be estimated and $H_0: \theta_1 = 0$ which is now equivalent to $H_0: \beta_1 + \beta_2 = 1$ can be tested using a t-test.

0004-14X1)

4. Since $\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2}$ (1)

and defining $\sum_{i=1}^n (x_i - \bar{x})^2$ as SST_x

we can rewrite (1) as:

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(\beta_0 + \beta_1 x_i + u_i)}{SST_x} \Rightarrow$$

$$\Rightarrow \hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})\beta_0 + \sum_{i=1}^n (x_i - \bar{x})\beta_1 x_i + \sum_{i=1}^n (x_i - \bar{x})u_i}{SST_x}$$

since $\sum_{i=1}^n (x_i - \bar{x}) = 0$ and $\sum_{i=1}^n (x_i - \bar{x})\beta_1 x_i = \beta_1 \cdot SST_x$

we now have $\hat{\beta}_1 = \beta_1 + \frac{\sum_{i=1}^n (x_i - \bar{x})u_i}{SST_x}$

so $E(\hat{\beta}_1) = \beta_1 + \frac{\sum_{i=1}^n [E(x_i - \bar{x})u_i]}{SST_x}$ since $E(\beta_1) = \beta_1$ and $E(SST_x) = SST_x$

Since $E(u_i | x_i) = 0$ we can write this as

$$E(\hat{\beta}_1) = \beta_1 + \frac{\sum_{i=1}^n (x_i - \bar{x})E(u_i)}{SST_x}$$

and since $E(u_i) = 0$ we are left with $E(\hat{\beta}_1) = \beta_1$